

AD A111323

**the
ohio
state
university**

research foundation

1314 kinnear road
columbus, ohio
43212

Report Title:

UNIFORM IMPROVEMENTS ON THE CERTAINTY EQUIVALENT RULE
IN A STATISTICAL CONTROL PROBLEM

by

L. Mark Berliner

Project Title:

PROBLEMS OF STATISTICAL INFERENCE WITH SPECIAL APPLICATIONS

Jagdish S. Rustagi, P.I.

DEPARTMENT OF THE NAVY
Office of Naval Research
Arlington, Virginia 22217

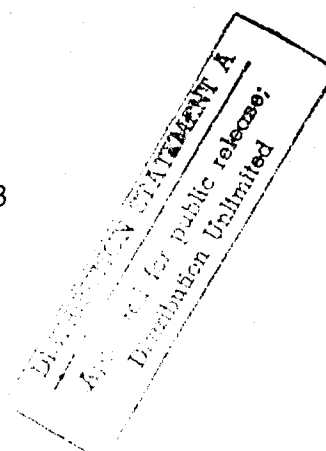
Contract No. N00014-78-C-0543

December, 1981

RF Project 761219/711308
Technical Report

FEB 24 1982

A



AD A111323-FILE COPY

Unclassified

SECURITY CLASSIFICATION OF THIS PAGE (When Data Entered)

| REPORT DOCUMENTATION PAGE | | READ INSTRUCTIONS BEFORE COMPLETING FORM | |
|---|-----------------------|--|--|
| 1. REPORT NUMBER | 2. GOVT ACCESSION NO. | 3. RECIPIENT'S CATALOG NUMBER | |
| | AD-A111 323 | | |
| 4. TITLE (and Subtitle) Uniform Improvements on the Certainty Equivalent Rule in a Statistical Control Problem | | 5. TYPE OF REPORT & PERIOD COVERED Technical Report | |
| | | 6. PERFORMING ORG. REPORT NUMBER 237 761219/711308 | |
| 7. AUTHOR(s) L. Mark Berliner | | 8. CONTRACT OR GRANT NUMBER(s) N000-14-78-C-0543 | |
| 9. PERFORMING ORGANIZATION NAME AND ADDRESS Dept. of Statistics Ohio State University 1958 Neil Ave., Columbus, Ohio 43210 | | 10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS NR 042-403 | |
| 11. CONTROLLING OFFICE NAME AND ADDRESS Office of Naval Research Department of Navy Arlington, Virginia 22207 | | 12. REPORT DATE December, 1981 | |
| | | 13. NUMBER OF PAGES 12 | |
| 14. MONITORING AGENCY NAME & ADDRESS (if different from Controlling Office) | | 15. SECURITY CLASS. (of this report) Unclassified | |
| | | 15a. DECLASSIFICATION/DOWNGRADING SCHEDULE | |
| 16. DISTRIBUTION STATEMENT (of this Report) Approved for public release; distribution unlimited. | | | |
| 17. DISTRIBUTION STATEMENT (of the abstract entered in Block 20, if different from Report) | | | |
| 18. SUPPLEMENTARY NOTES | | | |
| 19. KEY WORDS (Continue on reverse side if necessary and identify by block number) Control problem, differential inequalities, inadmissibility, multivariate normal distribution. | | | |
| 20. ABSTRACT (Continue on reverse side if necessary and identify by block number) The problem considered arises when it is desired to choose values for design variables in a linear model so that the resulting dependent random variable will be close to a prescribed constant. A decision theoretic analysis is considered. Control procedures which dominate, in risk, the commonly used certainty equivalent rule are presented. | | | |

DD FORM 1 JAN 73 1473

EDITION OF 1 NOV 69 IS OBSOLETE
S/N 0102-LF-014-6601

Unclassified

SECURITY CLASSIFICATION OF THIS PAGE (When Data Entered)

UNIFORM IMPROVEMENTS ON THE
CERTAINTY EQUIVALENT RULE IN A
STATISTICAL CONTROL PROBLEM¹

by

L. MARK BERLINER

¹Research supported in part by the National Science Foundation under Grant MSC-7802300 and by the Office of Naval Research under Contract N00014-78-C-0543(NR042-403).

Technical Report No. 237
Ohio State University
Department of Statistics
Columbus, Ohio 43210

December, 1981



A

UNIFORM IMPROVEMENTS ON THE CERTAINTY EQUIVALENT RULE IN A STATISTICAL CONTROL PROBLEM¹

L. Mark Berliner
Department of Statistics
Ohio State University
Columbus, Ohio, U.S.A.

1. The Control Problem. The problem considered here arises when it is desired to choose values for design variables in a linear model so that the resulting dependent random variable will be close to a prescribed constant. For discussions and applications of this problem, see Aoki [1], Dunsmore [6], Lindley [7], and Zellner [12].

The analysis here concerns the following transformed version of the problem as given in Basu [2] and Zaman [11]. Consider the linear model

$$y = z^t \beta + \epsilon,$$

where β is an unknown p -vector, z is an arbitrary p -vector, and $\epsilon \sim N(0, \sigma^2)$; σ^2 , possibly unknown. Furthermore, an estimate (independent of ϵ) of β , say $\hat{\beta}$, is available. Assume that $\hat{\beta} \sim N_p(\beta, \Lambda)$, where Λ is a known, positive definite matrix. The goal is to choose a controller $z_c(\hat{\beta})$ which performs well with respect to a control risk function R_c given by

$$\begin{aligned} (1.1) \quad R_c(z_c, \beta) &= E(y - y^*)^2 \\ &= E([z_c(\hat{\beta})]^t \beta + \epsilon - y^*)^2, \end{aligned}$$

where y^* is a non-zero constant, and the expectation is taken over both $\hat{\beta}$ and ϵ . Computation in (1.1) yields

$$(1.2) \quad R_c(z_c, \beta) = \sigma^2 + (y^*)^2 E[(y^*)^{-1} [z_c(\hat{\beta})]^t \beta - 1]^2.$$

¹Research supported in part by the National Science Foundation under Grant MSC-7802300 and by the Office of Naval Research under Contract N00014-78-C-0543(NR042-403).

Now, consider the following transformation:

$$\begin{aligned}x &= \Lambda^{-1/2} \hat{\beta} \\ \theta &= \Lambda^{-1/2} \beta \\ \delta &= (y^*)^{-1} \Lambda^{-1/2} z_c(\hat{\beta}).\end{aligned}$$

Application of this transformation in (1.2) yields the equivalent problem of choosing a decision rule $\delta(x) = (\delta_1(x), \dots, \delta_p(x))^t$, based on an observation $X \sim N_p(\theta, I)$, θ unknown, subject to incurring a loss,

$$(1.3) \quad L(\delta, \theta) = (\theta^t \delta - 1)^2.$$

This version of the control problem will be considered below.

The approach here is decision theoretic; control rules are evaluated in terms of their risk (expected loss) functions, $R(\delta, \theta) = E_\theta L(\delta(X), \theta)$. Our interest is the proposition of rules which dominate, in risk, the rule $\delta_m(x) = |x|^{-2} x$. ($|x|^2 = \sum_{i=1}^p x_i^2$.) This rule is the certainty equivalent controller. For brevity, discussion of the "certainty equivalent principle" (also known as the "separation principle") is omitted. The reader is referred to Aoki [1], Basu [2], and Zellner [12] for such discussions. However, in the spirit of this principle, note that the loss function implies that δ should be a estimator, in some sense, of the quantity $|\theta|^{-2} \theta$. Clearly, δ_m is the maximum likelihood estimator of this quantity.

As will be seen, a theorem of Zaman [11] implies that δ_m is inadmissible in all dimensions. (In fact, note that δ_m has infinite risk when p is 1 or 2.)

2. Previous Results. Most previous results are concerned with spherically symmetric (s.s.) rules, i.e., rules of the form $\delta(x) = \psi(|x|)x$. Many inadmissibility results are based on the asymptotic behavior of s.s. rules as $|x| \rightarrow \infty$. In particular, suppose that for some constant c ,

$$(2.1) \quad \delta(x) \propto (|x|^2 + c)^{-1} x$$

for $|x|$ sufficiently large. The essence of the available results is that if $c > 5 - p$, then δ is inadmissible. (See Zaman [10] and Berger, Berliner, and Zaman [4] for precise theorems.) Related work of Zaman [10] and Takeuchi [9] also concerns rules of the form (2.1). In particular, Zaman [10] showed that the value $c = 5 - p$ is asymptotically optimal (i.e., for all $|\theta|$ sufficiently large, $\delta(x) \propto (|x|^2 + 5 - p)^{-1} x$ has the smallest risk of all rules of the form (2.1)). Berger, Berliner, and Zaman [5] studied the admissibility of s.s. generalized Bayes controllers. Of special interest in this paper, it was shown that, under suitable regularity conditions on generalized prior measures, the corresponding generalized Bayes rules, that are given by $\delta(x) \propto (|x|^2 + 5 - p)^{-1} x$ for $|x|$ large, are admissible.

Three other results are directly applied in the discussion below. These results are paraphrased here. The reader is referred to the indicated references for precise statements.

RESULT 1. (Zaman [11]). If $\delta(x) = \Psi(|x|)x$ is admissible, then

- i) $0 \leq \Psi(r) \leq 1$.
- ii) $\lim_{r \rightarrow 0} \Psi(r) = 1$.
- iii) $\Psi(r)$ is a $C^{(\infty)}$ - function.

Note that Result 1; i), ii), imply that δ_m is inadmissible.

In Result 2, and the rest of the paper, the following notation is used:

DEFINITION: For any function Ψ , let

$$T_1(\Psi(r)) = \min \{1; \Psi(r)\}.$$

Result 1 can be viewed as a partial motivation of Result 2.

RESULT 2. (Zaman [11].) Let $\delta_0(x) = \psi_0(|x|)x$. If $P_\theta(\psi_0(|X|) > 1) > 0$, then δ given by $\delta(x) = T_1(\psi_0(|x|))x$ dominates δ_0 .

The next domination result was given by Berliner [5]. The derivation of this result is based on an integration by parts technique for risk analysis first introduced by Stein [8]. See Berliner [5] for a discussion and recent references.

For any differentiable function $\psi(r)$ let

$$\psi'(r) = \frac{d\psi(r)}{dr}.$$

Also, for all real η define the quantities $\xi(\psi)$ and $\zeta(\psi)$ by

$$\xi(\psi) = \int_0^\infty \left| \frac{d}{dr} [r^{(p+1)} \psi^2(r) \exp(-\frac{1}{2}r^2)] \right| \exp(r\eta) dr$$

and

$$\zeta(\psi) = [r^{(p+1)} \psi^2(r) \exp(-\frac{1}{2}r^2) \exp(r\eta)]_0^\infty.$$

RESULT 3. (Berliner [5]). Let $\delta(x) = \psi(|x|)x$ and $\delta_0(x) = \psi_0(|x|)x$. Suppose that both ψ and ψ_0 are continuous, piece-wise differentiable functions on $(0, \infty)$ such that, for all real η , the following conditions hold: (i) $\xi(\psi) < \infty$, (ii) $\xi(\psi_0) < \infty$, (iii) $|\zeta(\psi)| < \infty$, and (iv) $|\zeta(\psi_0)| < \infty$. If

$$(2.2) \quad \begin{aligned} \psi(r)\{2r\psi'(r) + (p+1-r^2)\psi(r) + 2\} &\geq \\ \psi_0(r)\{2r\psi_0'(r) + (p+1-r^2)\psi_0(r) + 2\} \end{aligned}$$

for all $r > 0$, and with strict inequality on a set of positive Lebesgue measure, then δ dominates δ_0 .

3. Main Results. Theorems 1 and 2 below present classes of procedures which dominate δ_m . These theorems are direct applications of Result 3.

THEOREM 1. Assume $p \geq 6$. Let $\delta(x) = \Psi(|x|)x$ where $\Psi(r) = (r^2 - g(r))^{-1}$.

Suppose that

- i) $g(r)$ is continuous, piece-wise differentiable, and non-decreasing,
- ii) $g(r)$ is not identically zero,
- and iii) $0 \leq g(r) \leq \min\{2(p-5); h(r^2, p)\}$, where $h(r^2, p)$ is given by

$$h(r^2, p) = \frac{r^2[r^2 + 3(p-3) - ((r^2 + 11 - p)^2 + 32(p-5))^{1/2}]}{2(r^2 + p - 3)}.$$

Then δ dominates δ_m .

Proof: First, note that δ_m clearly satisfies the requirements of Result 3 (when $p \geq 3$). It is also easy to verify that δ satisfies these requirements.

Step 1. To apply Result 3, let $\Psi_0(r) = r^{-2}$ and $\Psi(r) = (r^2 - g(r))^{-1}$. Computing (2.2) in this case yields

$$(3.1) \quad (r^2 - g(r))^{-1} \{-2r(2r - g'(r))(r^2 - g(r))^{-2} + (p+1-r^2)(r^2 - g(r))^{-1} + 2\} \geq r^{-2}\{2r(-2r^{-3}) + (p+1-r^2)r^{-2} + 2\}.$$

Simplification of (3.1) yields

$$(3.2) \quad (r^2 - g(r))^{-3}\{2r(g'(r) - 2r) + [(p+1-r^2) + 2(r^2 - g(r))](r^2 - g(r))\} \geq r^{-2}[(p-3)r^{-2} + 1].$$

It is easy to check that $g(r) \leq h(r^2, p)$ implies that $g(r) < r^2$ for all $r^2 > 0$. Using this fact, an algebraic manipulation of (3.2) implies the equivalent inequality

$$(3.3) \quad 2rg'(r) + r^{-4}g(r)\{[2(p-5)-g(r)]r^4 + [g(r)-3(p-3)]g(r)r^2 + (p-3)g^2(r)\} \geq 0.$$

Since $g'(r) \geq 0$ it is sufficient to verify that

$$(3.4) \quad [2(p-5)-g(r)]r^4 + [g(r)-3(p-3)]g(r)r^2 + (p-3)g^2(r) \geq 0.$$

Step 2: Let $Q(g)$ denote the L.H.S. of (3.4). Note that $Q(g)$ can be written as (suppressing the dependence of g on r)

$$(3.5) \quad Q(g) = (r^2+p-3)g^2 - r^2(r^2+3(p-3))g + 2(p-5)r^4.$$

Clearly $Q(g)$ is an upward opening quadratic function of g . Consider the smallest root, say \tilde{g} , of $Q(g)$. The quadratic formula implies that $\tilde{g} = h(r^2, p)$. Next, inspection of $h(r^2, p)$ implies that $h(r^2, p) > 0$ when $r^2 > 0$ (and, of course, $p \geq 6$). Hence, $Q(g) \geq 0$ for all $g \leq \tilde{g}$.

Finally, the dominance assertion then follows directly. ||

The remainder of this section is devoted to the development of a readily applicable version of Theorem 1. The motivation of this discussion is twofold. First, the function $h(r^2, p)$ in Condition iii) of Theorem 1 is rather cumbersome. Second, the specific g functions to be discussed in Section 4 are easily bounded by linear functions of r^2 .

THEOREM 2. Assume $p \geq 6$. Let δ be defined as in Theorem 1. Suppose that the corresponding g function satisfies Conditions i) and ii) of Theorem 1. If

$$0 \leq g(r) \leq (p-5)\min\{1, r^2/y^*\}$$

where $y^* > 0$ satisfies $h(y^*, p) = p-5$, then δ dominates δ_m .

The proof is based on the following lemma.

LEMMA 1. For $p \geq 6$ and $y \geq 0$, $h(y,p)$ is an increasing, concave function of y .

The proof of Lemma 1 is a straightforward differentiation argument and is omitted.

Proof of Theorem 2: By Theorem 1 it is sufficient to verify that $g(r) \leq \min\{(p-5); h(r^2,p)\}$.

First, by Lemma 1, $h(r^2,p) \leq p-5$ iff $r^2 \leq y^*$. Now consider a graph of $h(r^2,p)$ as a function of r^2 . The formula for the line connecting the origin and the point at which $h(r^2,p) = p-5$ is $[(p-5)/y^*]r^2$. Therefore, by concavity, $[(p-5)/y^*]r^2 \leq h(r^2,p)$ for all $r^2 \leq y^*$. ||

We close this section with a few remarks. First, the limit as $r^2 \rightarrow \infty$ of $h(r^2,p)$ is $3(p-3)/2$. Lemma 1 implies that if $p > 11$, Condition iii) of Theorem 1 simply reduces to $0 \leq g(r) \leq h(r^2,p)$. (Since $3(p-3)/2 < 2(p-5)$ when $p > 11$.)

Second, Theorem 2 can be generalized in the following sense. Theorem 2 is an application of Theorem 1 where the implicit upper bound on g is lowered from $2(p-5)$ to $p-5$. The same analysis could be performed for other upper bounds (moderated by the remark immediately above) if desired. However, Theorem 2 does include the important asymptotically optimal case.

Finally, the computation of y^* is required. For convenience, a partial list of the values of y^* (computed numerically) is given in Table 1.

Table 1. Selected Values of y^*

| p | y^* | p | y^* |
|----|-----------|----|-----------|
| 6 | 7.605551 | 15 | 20.000000 |
| 7 | 9.123106 | 16 | 21.280110 |
| 8 | 10.582576 | 17 | 22.549834 |
| 9 | 12.000000 | 18 | 23.810250 |
| 10 | 13.385165 | 19 | 25.062258 |
| 11 | 14.744563 | 20 | 26.306624 |
| 12 | 16.082763 | 25 | 32.433981 |
| 13 | 17.403124 | 30 | 38.440307 |
| 14 | 18.708204 | 40 | 50.206556 |

4. Proposed Control Rules. First, the following limitation in the application of Result 3 is noted. Substitution of $\Psi_0 = \Psi_m$ in (2.2) forces the R.H.S. of the inequality to be (positive) unbounded as $r \rightarrow 0$ (see (3.2)). Meanwhile, the L.H.S. of (2.2) remains bounded as $r \rightarrow 0$ for any admissible (or nearly admissible, in the sense of Result 1; i), ii)) rule. Hence, the differential inequality (2.2) cannot yield admissible alternatives to δ_m . In the case of Theorems 1 and 2, this fact is reflected in the implicit restriction that the Ψ functions satisfying the conditions of these theorems are bounded from below by $K|x|^{-2}$, for some constant K , as $|x| \rightarrow 0$. However, the combination of Result 2 and these theorems leads to reasonable alternatives to δ_m for $p \geq 6$.

PROPOSITION 1. Assume $p \geq 6$. Let $\delta(x) = \Psi(|x|)x$. Suppose δ satisfies the assumptions of Theorem 1 (or 2). Then δ_T given by

$$\delta_T(x) = T_1(\Psi(|x|))x$$

dominates δ_m .

Proof: Obvious. ||

REMARK: Unfortunately, the rules δ_T are also inadmissible since they violate the smoothness requirement of Result 1.

The final point of this discussion is the suggestion of functions $g(|x|)$ for actual use in the application of Theorem 2 and Proposition 1. The functions described below arise naturally as parts of a certain class of generalized Bayes rules. The reader is referred to Berger [3] and Berliner [5] for discussions. Only the required facts are given here.

Let $n \geq 1/2$. For $v > 0$ define the function $r_n(v)$ by

$$r_n(v) = \frac{\int_0^1 \lambda^n \exp(-\frac{1}{2}v\lambda) d\lambda}{\int_0^1 \lambda^{(n-1)} \exp(-\frac{1}{2}v\lambda) d\lambda}.$$

The following facts concerning r_n are needed here:

LEMMA 2. (Berger [3].) If $n \geq 1/2$, then

- i) $0 < r_n(v) < 2n$.
- ii) $r'_n(v) \geq 0$.
- iii) $\lim_{v \rightarrow \infty} r_n(v) = 2n$.
- iv) $\lim_{v \rightarrow \infty} r_n(v)/v = n/(n+1)$.
- v) $r_n(v)/v \leq n/(n+1)$.

Now, let $v = a|x|^2$ for some constant $a > 0$. Next, for $p \geq 6$, define $r^*(v)$ by $r^*(v) = r_n(v)$ for $n = (p-5)/2$. Then, clearly, by Lemma 2, $r^*(a|x|^2) \rightarrow p-5$ as $|x|^2 \rightarrow \infty$. Hence, the rule $\delta(x) = (|x|^2 - r^*(a|x|^2))^{-1}x$ is asymptotically optimal.

Also, by Lemma 2, note that

$$r^*(a|x|^2) \leq [(p-5)/(p-3)]a|x|^2.$$

Then, to apply Theorem 2, we require that

$$a \leq (p-3)/y^*.$$

PROPOSITION 2. Assume $p \geq 6$. Let δ^* be given by

$$\delta^*(x) = T_1([|x|^2 - r^*(a|x|^2)]^{-1})x.$$

Then for any constant a such that $0 < a \leq (p-3)/y^*$, δ^* dominates δ_m .

Proof: The proof is a direct application of the above arguments, Lemma 2, and Proposition 1. ||

5. Comments. i) The rules δ^* proposed above display desirable properties:

- a) They are relatively easy to compute.
- b) Their behavior for $|x|$ large is similar to that of the generalized Bayes, admissible rules discussed in Berliner [5].
- c) They are asymptotically optimal.

However, they are not admissible since the smoothness requirement of Result 1 is violated.

ii) A common criterion for choosing among decision rules is minimaxity. We simply note here that δ_m is minimax when $p \geq 3$. Hence, δ^* is also minimax. See Berliner [5] for proofs and discussion.

iii) Another natural control procedure often considered is the uniform measure, generalized Bayes rule $\delta_u(x) = (1+|x|^2)^{-1}x$. This rule is admissible for $p \leq 4$, but inadmissible for $p \geq 5$. Several authors have shown that δ_u is dominated by δ_m when $p \geq 5$. Hence, the rules δ^*

proposed above also dominate δ_u when $p \geq 6$. For further discussion, references for the above results, and another class of rules which dominate δ_u , see Berliner [5].

Acknowledgments. I wish to express my appreciation to Jim Berger for his help and comments. My thanks also go to Jayne Sage for programming the computations for Table 1.

REFERENCES

- [1] Aoki, M. (1967). Optimization of Stochastic Systems. Academic Press, New York.
- [2] Basu, A. (1974). Control level determination in regression models. Technical Report 139, Economic Series, Institute for Mathematical Studies in the Social Sciences, Stanford University.
- [3] Berger, J. (1980). A robust generalized Bayes estimator and confidence region for a multivariate normal mean. Ann. Statist. 8, 716-761.
- [4] Berger, J., Berliner, L. M., and Zaman, A. (1981). General admissibility and inadmissibility results for estimation in a control problem. Mimeograph Series #79-34, Statistics Department, Purdue University.
- [5] Berliner, L. M. (1981). Improving on inadmissible estimators in the control problem. Technical Report 229, Department of Statistics, The Ohio State University.
- [6] Dunsmore, I. R. (1969). Regulation and optimization. J. Roy. Statist. Soc. Ser. B 31, 160-170.
- [7] Lindley, D. V. (1968). The choice of variables in multiple regression. J. Roy. Statist. Soc. Ser. B 30, 31-66.
- [8] Stein, C. (1973). Estimation of a mean of a multivariate distribution. Proc. Prague Symp. Asymptotic Statist., 345-381.
- [9] Takeuchi, K. (1968). On the problem of fixing the level of independent variables in a linear regression function. IMM 367, Courant Institute of Math. Sciences, New York University.
- [10] Zaman, A. (1977). An application of Stein's method to the problem of single period control of regression models. Technical Report 231, Economic Series, Institute for Mathematical Studies in the Social Sciences, Stanford University.

- [11] Zaman, A. (1981). A complete class theorem for the control problem, and further results on admissibility and inadmissibility. Ann. Statist. 9, 812-821.
- [12] Zellner, A. (1971). An Introduction to Bayesian Inference in Econometrics. John Wiley, New York.